

**An 11-Factor Heath, Jarrow, and Morton Stochastic Volatility Model  
for the Government of Russia Yield Curve,  
Using Daily Data from January 4, 2003 through August 31, 2021  
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**ABSTRACT**

*Please note: Kamakura Corporation term structure models are updated monthly. For the most recent set of coefficients, contact [info@kamakuraco.com](mailto:info@kamakuraco.com)*

This paper analyzes the number and the nature of factors driving the movements in the Government of Russia yield curve from January 4, 2003 through August 31, 2021. The process of model implementation reveals a number of important insights for interest rate modeling generally. First, model validation of historical yields is important because those yields are the product of a third-party curve fitting process that may produce spurious indications of interest rate volatility. Second, quantitative measures of smoothness and international comparisons of smoothness provide a basis for measuring the quality of simulated yield curves. Third, we outline a process for incorporating insights from the Japanese and European experience with negative interest rates into term structure models with stochastic volatility in Russia and other countries. Fourth, we compare data availability for Russia with broad international experience to measure the risk that a simulation beyond historical rate levels in Russia could go awry. Finally, we illustrate the process for comparing stochastic volatility and affine models of the term structure. We conclude that stochastic volatility models have a superior fit to the history of yield movements in the Government of Russia market.

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We also recommend that Government of Russia interest rate risk analysis employ the full “World” 13-country term structure model rather than relying solely on Russia data alone.

## **A 11-Factor Heath, Jarrow, and Morton Model for the Government of Russia Yield Curve, Using Daily Data from January 4, 2003 through August 31, 2021**

Government yield curves are a critical input to the risk management calculations of central banks, bank regulators, major banks, insurance firms, fund managers, pension funds, and endowments around the world. With the internationalization of fixed income investing, it is important to understand the dynamics of movements in yield curves worldwide, in addition to the major bond markets like those in Frankfurt, London, New York and Tokyo. In this paper, we fit a multi-factor Heath, Jarrow and Morton model to daily data from the Government of Russia securities market over the period from January 4, 2003 to August 31, 2021. The modeling process reveals a number of important implications for term structure modeling in other government bond markets.

Section I discusses the origin and characteristics of the daily data base of Government of Russia yields provided by the Central Bank of Russia. We discuss yield curve smoothness and volatility fitting as measures for judging the quality of government-generated yield curve time series, including Government of Russia yield curves. We conclude that the Central Bank of Russia time series is realistically smooth and a reliable foundation for term structure modeling. If anything, the data is too smooth and shows so few humps in the yield curve that we presume a specific functional form was imposed during the smoothing process. This compares with recent findings from Australia, Japan, and Thailand where we found that yield curves were unrealistically rough and that modification of the input data was necessary for a realistic model, a standard part of a Bayesian model validation process.

Section II outlines the process for determining whether the interest rate volatility for the factors driving the Government of Russia yield curve is constant (i.e., non-random, an “affine” model) or stochastic, typically expressed as a function of the level of interest rates. We note the extensive experience with negative interest rates in the European and Japanese government bond markets and use insights from that experience in fitting volatility in the Government of Russia market. Section III describes the process of fitting five different Heath, Jarrow, and Morton models to Government of Russia yield data: models with 1, 2, 3, 6 and 11 factors. We conclude Section III with Bayesian model validation procedures based on stress tests of the derived volatility and drift functions. Section IV concludes the paper. Appendix A illustrates a sample model validation process for widely used one factor term structure models. Appendix B includes additional analysis of empirical drift and stress tests of rate volatility and compares Russia results to the 13-country World term structure model.

### **I. Government of Russia Data: Special Characteristics**

A multi-factor term structure model is the foundation for best practice asset and liability management, market risk, economic capital, interest rate risk in the banking book, stress-testing, and the internal capital adequacy assessment process. The objective in this paper is to illustrate the derivation of a multi-factor Heath Jarrow and Morton model of the Government of Russia yield curve. As a by-product, the analysis has the potential to detect common data problems associated with yield curve histories and employs a standard methodology for quantification and resolution of those problems. Previous

implementations of multi-factor Heath, Jarrow and Morton models have covered the following bond market sectors:

Australia	Commonwealth Government Securities
Canada	Government of Canada Securities
France	French Government Securities
Germany	German Bunds
Italy	Italian Government Securities
Japan	Japanese Government Bonds
Russia	Russia Government Securities
Singapore	Singapore Government Securities
Spain	Spanish Government Bonds
Sweden	Swedish Government Securities
Thailand	Thai Government Securities
United Kingdom	United Kingdom Government Bonds
United States	Treasury Securities

The first step in data model validation for the Government of Russia market is to examine the historical availability of bond yields over time. This availability is summarized in Table I.

**Table I**

Kamakura Corporation  
HJM 11 Factor Model  
Russia Government Securities  
Using Daily Data from January 4, 2003 through August 31, 2021  
Date of Analysis: August 31, 2021  
Number of Observations by Data Regime

Data Regime	Start Date	End Date	Number of Observations
rus: Yields to 30 Years	1/8/2003	8/27/2021	4,626
Total Observations			4,626

The data shows that the Government of Russia’s data history is atypical in that the Central Bank of Russia time series shows no changes in “data regime,” i.e., which of the maturities are available on a given date. Canada, where the Bank of Canada is the data source, is another rare example of a single data regime.

Because our Heath, Jarrow and Morton analysis makes use of a yield curve with 91-day (quarterly) forward rate segments, the next step in data model validation is to fit quarterly forward rates to the raw coupon-bearing bond yields. The smoothness of the resulting forward rates will be a function of both the quality of the raw data from a smoothness point of view and the smoothness implied by the secondary smoothing process. To ensure the maximum smoothness from the secondary smoothing process, we use the maximum smoothness forward rate methodology of Adams and van Deventer [1994], as corrected in van Deventer and Imai [1996]. Adams and van Deventer show that the maximum smoothness method overcomes the problems of the cubic spline approach of McCulloch, and, unlike the Svensson [1994] approach, allows for a perfect fit to the raw data provided by the [Central Bank of Russia](#). See Jarrow [2014] for information on the problems with Svensson yield curve fitting.

We then conduct a visual inspection of the resulting forward rates implied by the raw data. The smoothness of the quarterly forward rate curve can be measured quantitatively using the quarterly forward rates implied by the Government of Russia yield curves. For a yield curve that consists of  $N$  quarterly forward rates, the discrete smoothness statistic at time  $t$ ,  $Z_N(t)$ , is the sum of the squared second differences in the forward rates, as explained by Adams and van Deventer [1994]. A closed form continuous smoothing statistic can also be calculated when the functional form of the continuous forward rate is known (as it is in this case). The discrete smoothness statistic is given here:

$$Z_N(t) = \sum_{i=3}^N [(f_i(t) - f_{i-1}(t)) - (f_{i-1}(t) - f_{i-2}(t))]^2$$

The Government of Russia series is realistically smooth, and we use that data with confidence in what follows. We apply a final screen for outliers in the process of deriving the stochastic volatility functions. Yield curves on the following days were omitted from estimation because they implied implausible forward rate shocks in the third and fifth 91-day forward rates:

April 29, 2004

May 6, 2004

May 11, 2004

May 12, 2004

May 13, 2004

May 14, 2004

May 17, 2004

May 18, 2004

May 20, 2004

May 21, 2004

May 25, 2004

May 26, 2004

May 27, 2004

May 28, 2004

May 31, 2004

June 1, 2004

June 3, 2004

June 4, 2004

June 8, 2004

June 9, 2004

June 10, 2004

June 11, 2004

June 15, 2004

June 16, 2004

June 17, 2004

June 18, 2004

June 21, 2004

June 22, 2004

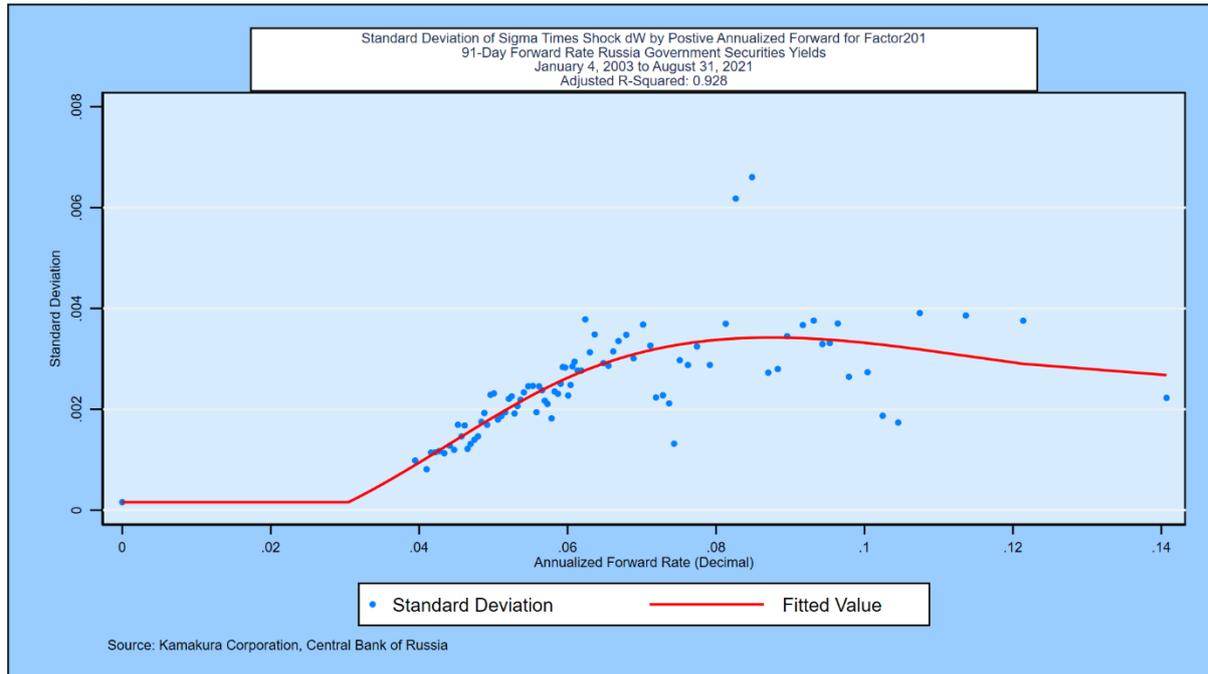
June 23, 2004  
June 24, 2004  
June 25, 2004  
June 28, 2004  
June 29, 2004

July 1, 2004  
July 2, 2004  
July 5, 2004  
July 6, 2004  
July 7, 2004  
July 8, 2004  
July 9, 2004  
July 12, 2004  
July 13, 2004  
July 14, 2004  
July 15, 2004  
July 19, 2004

## **II. Constant versus Stochastic Volatility**

Constant volatility (“affine”) term structure models are commonly used for their ease of simulation and estimation of “future expected rates” in order to determine the “term premium” in current yields. Prominent examples are Adrian, Crump and Moench [2013], Kim and Wright [2005], and Duffie and Kan [1996]. On the other hand, the weight of the empirical evidence in most of the countries studied to date indicates that interest rate volatility does vary by the level of the corresponding forward rate. To illustrate that fact, we studied the shortest forward rate on the Government of Russia curve on a daily basis from January 4, 2003 through August 31, 2021. We ordered the data from lowest forward rate level to highest forward rate level. We formed non-overlapping groups using the larger of 50 observations or 1/100<sup>th</sup> of total observations each and calculated both the standard deviation of 91-day forward rate changes and the mean beginning-of-period forward rate in each group. The results are plotted in Exhibit III:

### Exhibit III

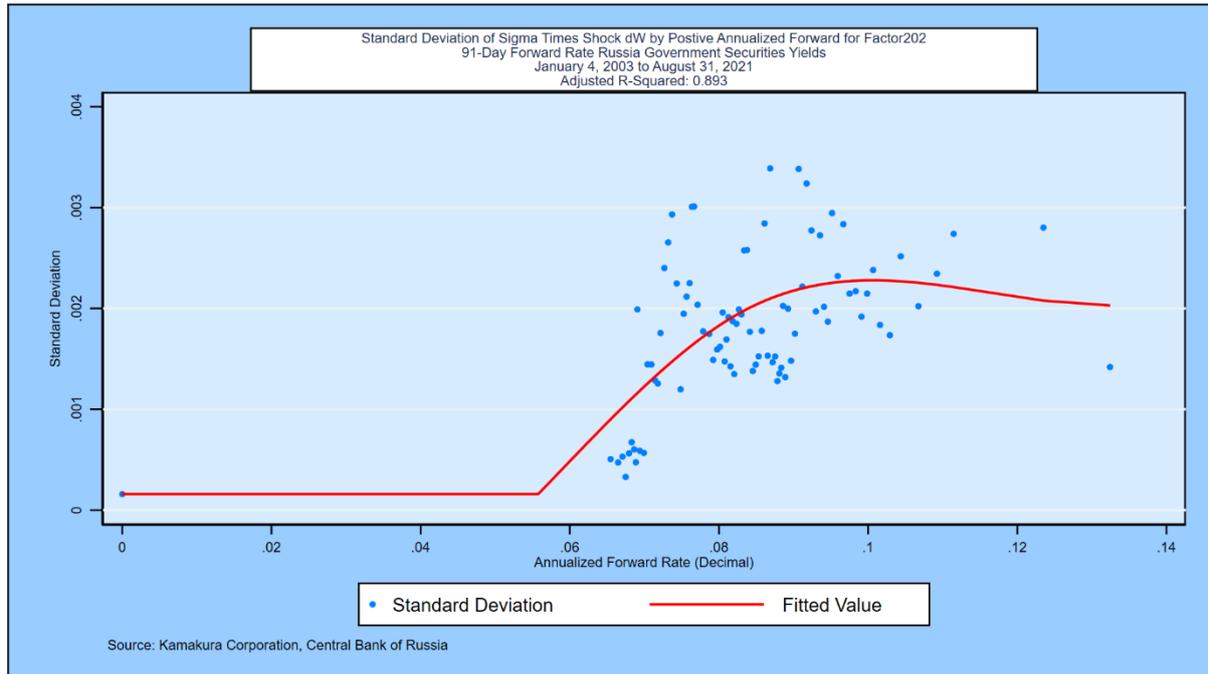


A 2-segment cubic spline of annualized forward rates explains 92.8%<sup>2</sup> of the variation in the standard deviation of forward rate changes for these ordered groups. This is the volatility function used when extracting the first random factor driving the Government of Russia curve. Note that the right-hand side of the curve has been constrained to have a first derivative of zero at a high level of rates.<sup>3</sup> The rise in volatility in higher rate environments has been confirmed in the government securities markets for Australia, Canada, France, Germany, Italy, Japan, Russia, Singapore, Spain, Sweden, Thailand, the United Kingdom, and the United States. Exhibit IV shows the results for the second risk factor in the Government of Russia market, the idiosyncratic movements in the quarterly forward rate maturing in 10 years:

<sup>2</sup> The reported accuracy is affected by the large weighting placed on observations below or near zero to be consistent with low-rate experience in 13 countries.

<sup>3</sup> This constraint is one method for imposing the cap in stochastic volatilities suggested by Heath, Jarrow, and Morton [Econometrica, 1992] to prevent a positive possibility of (a) infinitely high rates or (more practically) (b) unrealistically high rates.

## Exhibit IV



The cubic spline stochastic volatility specification explains 89.3% of the observed variation in forward rate volatility in the quarterly forward rate maturing at the 10-year point on the Government of Russia yield curve. We have imposed the same constraint on the first derivative and require that the fitted volatility not be less than the lowest observed volatility, which we discuss later in this section. In both cases, there are no observations of volatility in the forward rate range from 0 to 5%, which is why we recommend later in the paper that the 13-country World model would add to the robustness of term structure model parameters for the Government of Russia HJM model.

Exhibit V shows the historical movements in Government of Russia zero-coupon yields over the historical period studied. In later sections, we compare the span of the Russia data set with the range of rates and maturities contained in the 13-country World model. The Russia data set spans 53.77% of the World experience, arrayed by rate level and years to maturity. This is of great importance and we return to this topic in depth below.

## Exhibit V

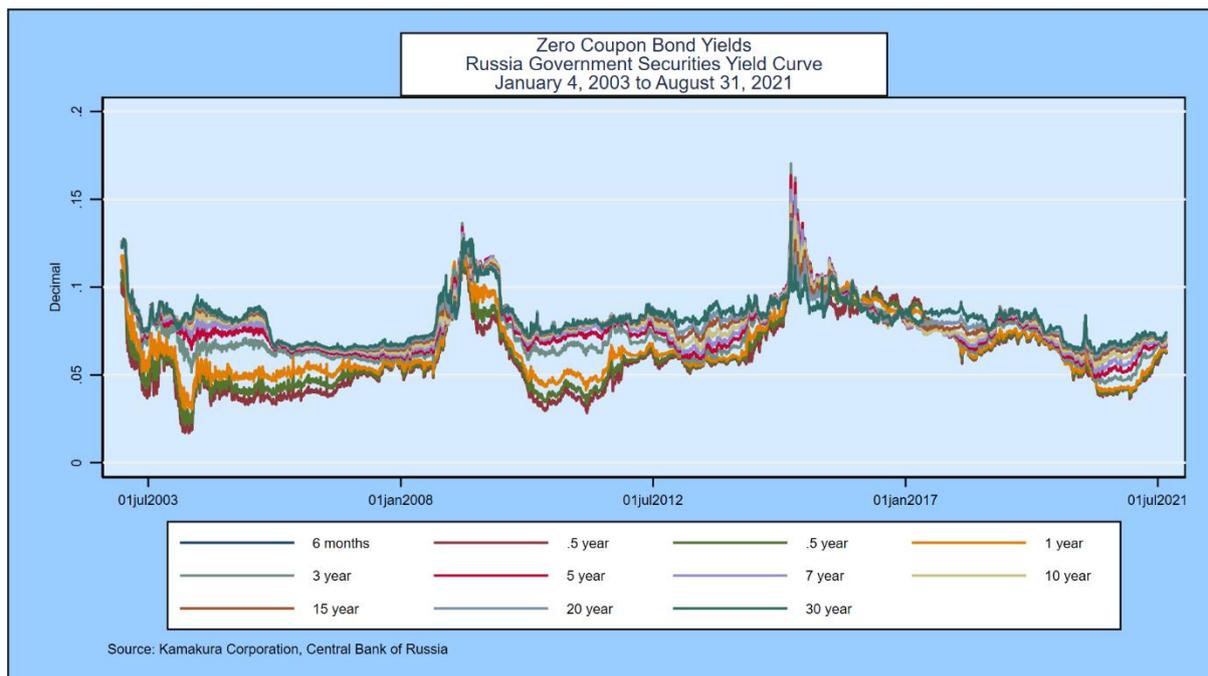
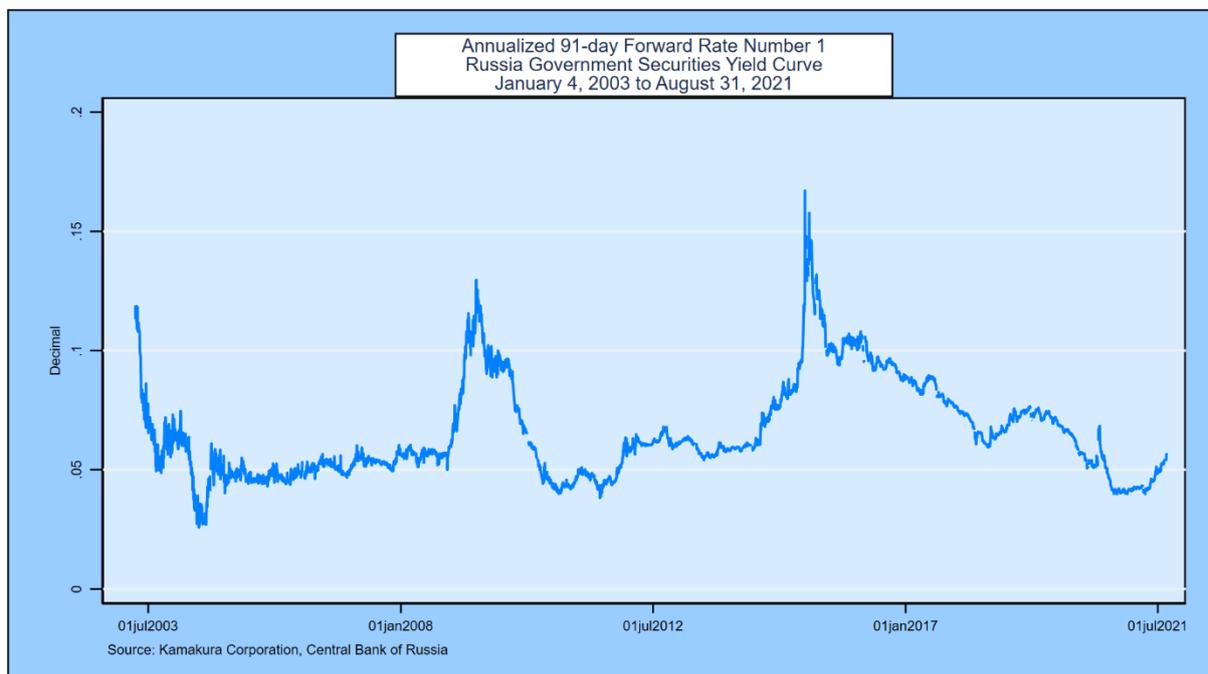


Exhibit VI below shows the evolution of the first quarterly forward rate (the forward that applies from the 91<sup>st</sup> day through the 182<sup>nd</sup> day) over the same time period:

## Exhibit VI



We use three statistical tests to determine whether or not the hypothesis of normality for forward rates and zero-coupon bond yields should be rejected at the 5% level: the [Shapiro-Wilk](#) test, the [Shapiro-Francia](#) test, and the [skew test](#), all of which are available in common statistical packages. The Russia data history is near but still under the

maximum number of observations allowed by these tests. The hypothesis of normality was rejected at each 91-day forward rate maturity date.

The rejection of the hypothesis of normality is problematic for constant (more precisely, non-random) coefficient or “affine” term structure models. In most of the other countries studied, the hypothesis of normality has been rejected strongly as well. Given these results, we proceed with caution on the implementation of the affine model.

In Chapter 3 of [Advanced Financial Management](#) (second edition, 2013), van Deventer, Imai and Mesler analyze the frequency with which forward rates move up together, down together or remain unchanged. This exercise informs the Heath, Jarrow and Morton parameter fitting process and is helpful for the model validation questions posed in the Appendix. We perform the yield curve shift analysis using 4,626 days of quarterly forward rates for the Government of Russia yield curve. We analyze the daily shifts in the forward rates on each business day from January 4, 2003 through August 31, 2021. The results are given in Table II:

**Table II**

**Kamakura Corporation**  
**HJM 11 Factor Model**  
**Russia Government Securities**  
**Using Daily Data from January 4, 2003 through August 31, 2021**  
**Date of Analysis: August 31, 2021**

<b>Type of Yield Shift</b>	<b>Number of Observations</b>	<b>Percent of Observations</b>
<b>All yields shift up</b>	659	14.25
<b>All yields shift down</b>	724	15.65
<b>All yields are unchanged</b>	3	0.06
<b>Yield curve twists</b>	3,240	70.04
<b>Total</b>	4,626	100.00

Kamakura Corporation, Central Bank of Russia

Yield curve shifts were all positive, all negative, or all zero 14.25%, 15.65%, and 0.06% of the time, a total of 29.96% of all business days. The predominant yield curve shift was a twist, with a mix of positive changes, negative changes, or zero changes. These figures are similar to those for the 12 other countries for which we estimate term structure model parameters on a regular basis. These twists, which happen 70.04% of the time in Russia, cannot be modeled accurately with the conventional implementation of one factor term structure models. The percentage of twists is somewhat lower than average for the other 12 countries in the world sample, which may indicate a specific functional form was imposed on raw data in the smoothing process.

Another important aspect of yield curves is the number of local minima and maxima that have occurred over the modeling period. The results for the Government of Russia market are given in Table III:

**Table III****Analysis of Number of Local Minima and Maxima Each Day**

<b>Number of Humps</b>	<b>Number of Observations</b>	<b>Percent of Observations</b>
<b>0 local minimum and maximum</b>	2,631	56.87
<b>1</b>	1,548	33.46
<b>2</b>	302	6.53
<b>3</b>	104	2.25
<b>4</b>	29	0.63
<b>5</b>	11	0.24
<b>6</b>	1	0.02
<b>7</b>	0	0.00
<b>8</b>	0	0.00
<b>9</b>	0	0.00
<b>10 or more</b>	0	0.00

Kamakura Corporation, Central Bank of Russia

The number of days with 0 or 1 humps (defined as the sum of local minima and maxima on that day's yield curve) was 90.33% of the total observations in the data set. Again, this is a much higher percentage than normal, so the evidence of a specific functional form used in smoothing is mounting.

Finally, before proceeding, we count the number of occurrences of negative rates for each forward rate segment of the yield curve over the history provided by the Bank of Russia and report on the observed 91-day volatility of forward rates when the start of the period annualized forward rate is negative, zero, and positive.

**Table IV**

Kamakura Corporation

HJM 11 Factor Model

Russia Government Securities

Using Daily Data from January 4, 2003 through August 31, 2021

Date of Analysis: August 31, 2021

Source: Kamakura Corporation, Central Bank of Russia

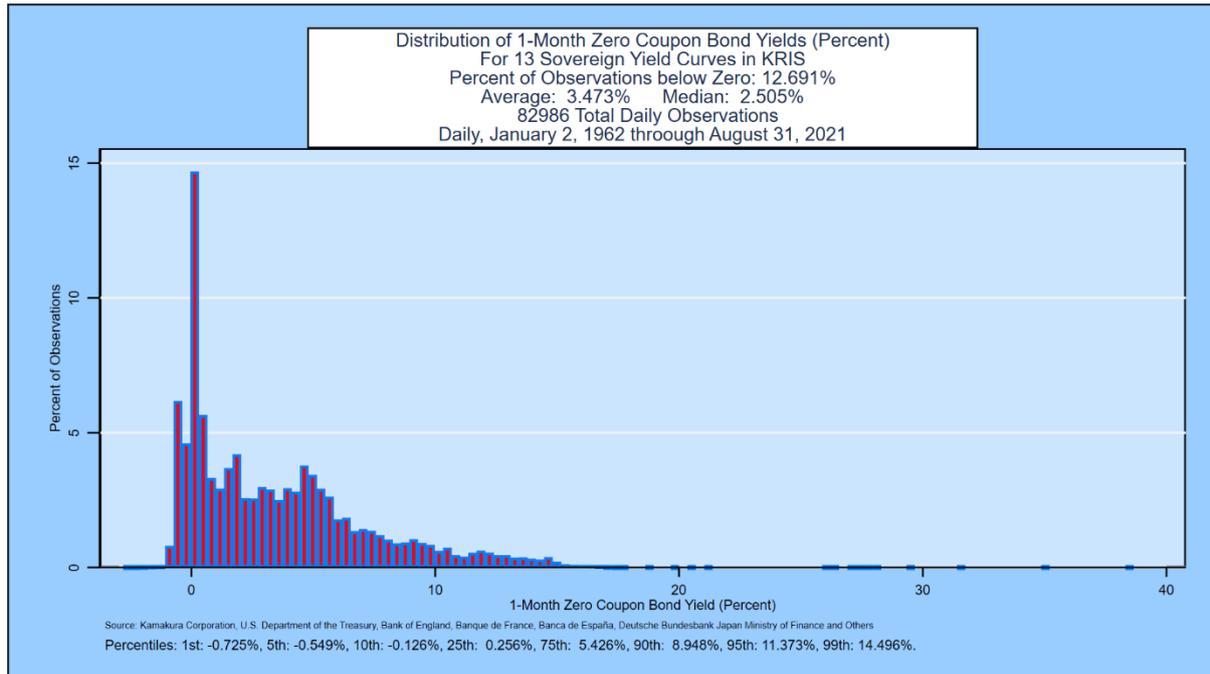
**Count of Negative 91-Day Forward Rate Forward Rates**

<b>Period Number</b>	<b>Observations</b>	<b>Negative</b>	<b>Zero</b>	<b>Positive</b>
1	4626	0	0	4626
2	4626	0	0	4626
3	4626	0	0	4626
4	4626	0	0	4626
5	4626	0	0	4626
6	4626	0	0	4626
7	4626	0	0	4626
8	4626	0	0	4626
9	4626	0	0	4626
10	4626	0	0	4626
11	4626	0	0	4626
12	4626	0	0	4626

The Government of Russia yield curves imply no negative yields over the first 12 91-day periods analyzed.

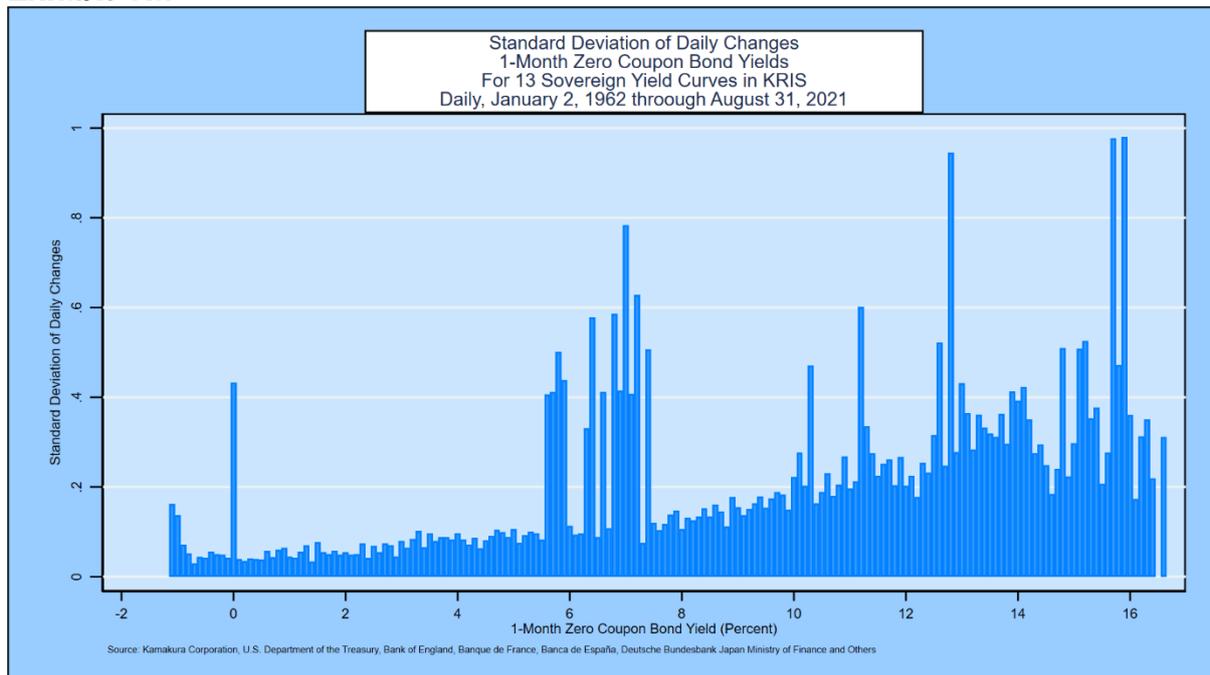
The next exhibit shows that negative rates at the 1-month time horizon have become common among the 13-country World interest rate data base maintained by Kamakura Corporation. Of 82,986 observations at a 1-month maturity, 12.691% of World experience has been negative.

## Exhibit VII



Observed volatility in 1-month zero-coupon bond yields as a function of interest rate level for the standard 13-country World data set is shown in Exhibit VII:

## Exhibit VIII



### III. Fitting Heath, Jarrow, and Morton Parameters

A simple first step in constructing a multi-factor Heath, Jarrow and Morton model is to conduct principal components analysis on the forward rates that make up the relevant yield curve. For the Government of Russia market, at its longest maturity, these

quarterly segments consist of one three-month spot rate and 119 forward rates. Over 4,329 observations, the principal components analysis indicates in Table V that the first factor explains 79.33% of the movement in forward rates over the full curve. This is a much higher percentage than average for the other 12 countries in the World model, which again implies a specific functional form was used in smoothing. For a high degree of explanatory power, the principal components analysis indicates that 6 to 11 factors will be necessary.

**Table V**

Principal components/correlation	Number of obs	=	4,329
	Number of comp.	=	15
	Trace	=	119
Rotation: (unrotated = principal)	Rho	=	1.0000

Component	Eigenvalue	Difference	Proportion	Cumulative
Comp1	94.4072	78.3131	0.7933	0.7933
Comp2	16.0941	10.2422	0.1352	0.9286
Comp3	5.85195	4.26383	0.0492	0.9778
Comp4	1.58812	.879854	0.0133	0.9911
Comp5	.708261	.484099	0.0060	0.9971
Comp6	.224161	.157128	0.0019	0.9989
Comp7	.0670336	.0357718	0.0006	0.9995
Comp8	.0312619	.0156971	0.0003	0.9998
Comp9	.0155648	.00643097	0.0001	0.9999
Comp10	.00913381	.00703407	0.0001	1.0000
Comp11	.00209975	.00129609	0.0000	1.0000
Comp12	.000803661	.000599253	0.0000	1.0000
Comp13	.000204408	.000127413	0.0000	1.0000
Comp14	.0000769949	.0000661333	0.0000	1.0000
Comp15	.0000108616	7.11492e-06	0.0000	1.0000
Comp16	3.74670e-06	3.74670e-06	0.0000	1.0000
Comp17	0	0	0.0000	1.0000
Comp18	0	0	0.0000	1.0000
Comp19	0	0	0.0000	1.0000
Comp20	0	0	0.0000	1.0000

With this analysis as background, we begin the Heath, Jarrow, and Morton fitting process.

In the studies done so far, the number of statistically significant factors is summarized below:

Australia:	Commonwealth Government Securities,	14 factors
Canada:	Government of Canada Securities,	7 factors
France:	French Government Bonds	7 factors
Germany:	Bunds,	14 factors
Italy:	Italian Government Bonds	11 factors
Japan:	Japanese Government Bonds	8 factors
Russia:	Russia Government Bonds (prior version)	11 factors
Singapore:	Singapore Government Securities	9 factors
Spain:	Spanish Government Securities	11 factors
Sweden:	Swedish Government Securities,	11 factors
Thailand:	Thai Government Securities	11 factors
United Kingdom:	Government Securities,	14 factors
United States:	U.S. Treasury Bonds	10 factors
World:	13 countries' government securities	12 factors

We now fit a multi-factor Heath, Jarrow and Morton model to Government of Russia zero-coupon yield data from January 4, 2003 to August 31, 2021. The procedures used to derive the parameters of a Heath, Jarrow and Morton model are described in detail in Jarrow and van Deventer (June 16, 2015 and May 5, 2017).

We followed these steps to estimate the parameters of the model:

- We extract the zero-coupon yields and zero-coupon bond prices for all quarterly maturities out to 30 years for all daily observations for which the 30-year zero-coupon yield is available. This is done using Kamakura Risk Manager, version 10.1, using the maximum smoothness forward rate approach to fill the quarterly maturity gaps in the zero-coupon bond data.
- We use overlapping 91-day intervals to measure changes in forward rates, avoiding the use of “quarterly” data because of the unequal lengths of calendar quarters. Because overlapping observations trigger auto-correlation, “HAC” (heteroscedasticity and autocorrelation consistent) standard errors are used. The methodology is that of Newey-West with 91-day lags.
- We consider 11 potential explanatory factors: the idiosyncratic portion of the movements in quarterly forward rates that mature at the 11 yield curve maturities given below.
- We calculate the discrete changes in forward returns as described in the parameter technical guide. Because the discrete changes are non-linear in the no-arbitrage framework of Heath, Jarrow, and Morton, we use generalized linear models and the proper HJM exponential functional form to fit interest rate volatility.
- We use a different non-linear regression for each segment of the yield curve. We considered a panel-based approach, but we rejected it for two reasons: first, the movement of parameters as maturity lengthens is complex and not easily predictable before estimation; second, the residual unexplained error in forward rates is very, very small, so the incremental merit of the panel approach is minimal.
- We then begin the process of creating the orthogonalized risk factors that drive interest rates using the Gram-Schmidt procedure. Without loss of generality, these factors are assumed to be uncorrelated independent

random variables that have a normal distribution with mean zero and standard deviation of 1.

- Because interest volatility is assumed to be stochastic, simulated out-of-sample forward rates will not in general be normally distributed. We also calculate constant volatility parameters and choose the most accurate from the constant volatility and stochastic volatility models estimated.
- In the estimation process, we added factors to the model as long as each new factor provided incremental explanatory power. The standard suite of models in both cases includes 1 factor, 2 factors, 3 factors, 6 factors and “all factors,” which varies by country.

We postulate that interest rate volatility for each forward rate maturity  $k$  is a cubic function of the annualized forward rate that prevails for the relevant risk factor  $j$  at the beginning of each 91-day period:

$$\sigma_{jk} = \max [b_{0,jk}, b_{0,jk} + b_{1,jk}f + b_{2,jk}f^2 + b_{3,jk}f^3] \text{ if } f > 0, \\ \sigma_{jk} = b_{0,jk} \text{ if } f \leq 0,$$

When the initial forward rate is negative, we postulate that interest rate volatility is a constant<sup>4</sup>. Using Japan volatility data reported above, the measured  $b_{0,jk}$  was 0.018553%. Using the 13-country World model, the value of  $b_{0,jk}$  was 0.0482%.

We use the resulting parameters and accuracy tests to address the hypothesis that a one factor model is “good enough” for modeling Government of Russia yields in Appendix A. We report the accuracy results for 1, 2, 3, 6 and all (11) factors. The factors are the idiosyncratic variation in quarterly forward rates at each of 11 maturities. The factors, described by the maturity of the 91-day forward rate used, are added to the model in this order:

#### **Data Regime 1 (Maturities of 30 years or less)**

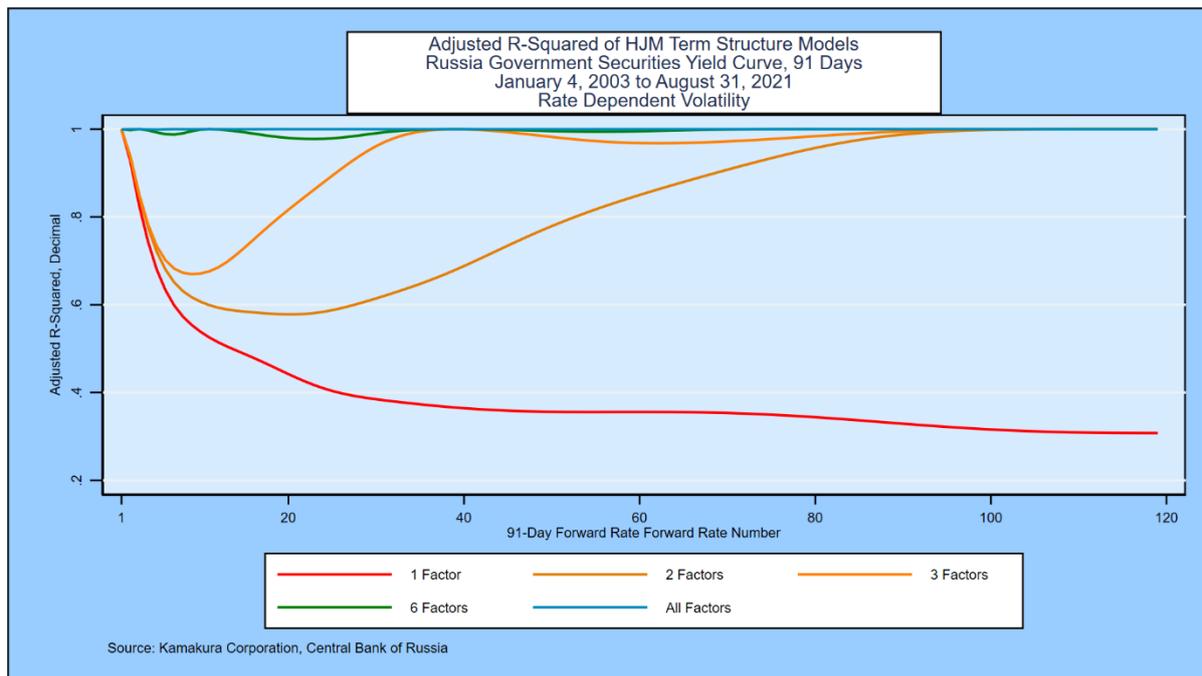
Factor 1:	6 months
Factor 2:	30 years
Factor 3:	10 years
Factor 4:	3 years
Factor 5:	20 years
Factor 6:	1 year
Factor 7:	7 years
Factor 8:	2 years
Factor 9:	5 years
Factor 10:	9 months
Factor 11:	15 years

Exhibit IX summarizes the adjusted r-squared for the non-linear equations for each of the 11 quarterly forward rate segments that make up the Government of Russia yield curve:

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<sup>4</sup> We make this assumption while waiting for more data on negative rates to accumulate. At the appropriate time, a more interesting functional form will be employed.

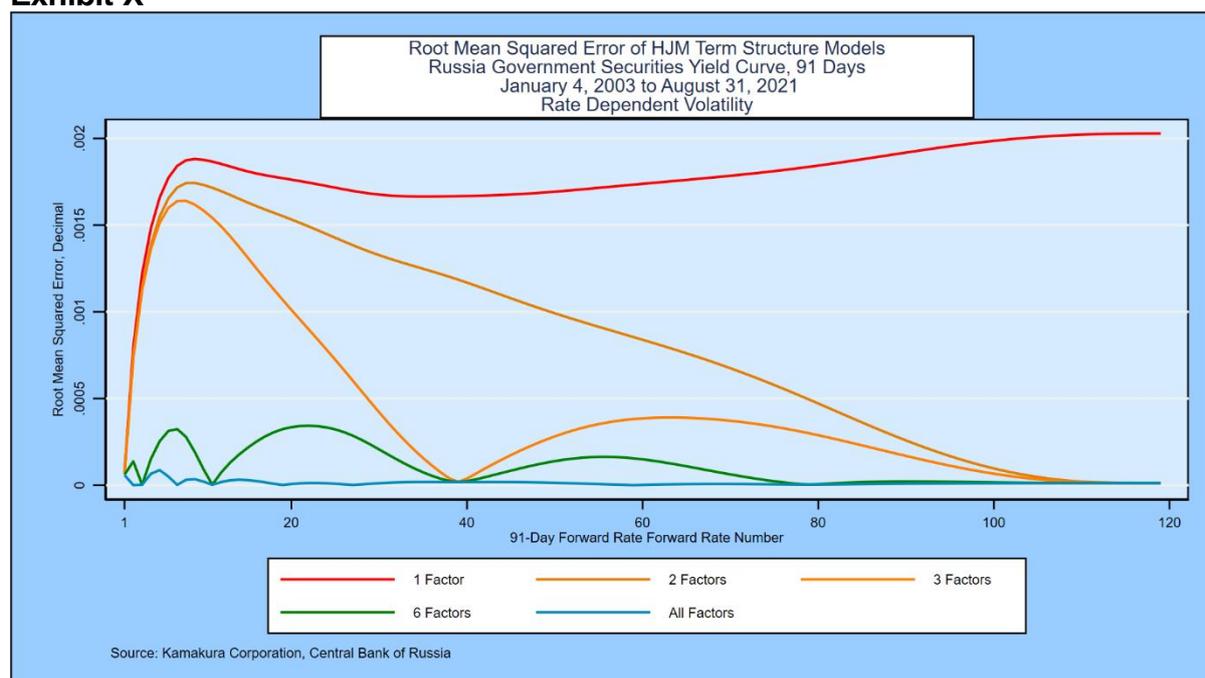
## Exhibit IX



The adjusted r-squared for the best practice model over each of the forward rates is plotted in blue and is near 100% for all 119 quarterly segments of the yield curve. The one factor model in red, by contrast, does a poor job of fitting 91-day movements in the quarterly forward rates. The adjusted r-squared is good, of course, for the first forward rate since the short rate is the standard risk factor in a one factor term structure model. Beyond the first quarter, however, the explanatory power of a one-factor model declines rapidly. The adjusted r-squared of the one-factor model never exceeds 40% after the first 24 quarterly forward rates and is below that level at most maturities.

The root mean squared error for the 1, 2, 3, 6 and all (11) factor stochastic volatility model is shown in Exhibit X.

## Exhibit X



The root mean squared error for the 11-Factor model is less than 0.02% at every maturity along the yield curve. This result should not come as a surprise to a serious analyst, because it is very similar to the results of the best practice Heath, Jarrow and Morton term structure models for the other 12 government bond markets studied.

### Bayesian Considerations in Model Validation

Kamakura term structure model validation is conducted in the spirit of Bayesian iterative model building as outlined by Gelman et al [2013]. This quote<sup>5</sup> from Gelman et al explains the Bayesian estimation process:

“The process of Bayesian data analysis can be idealized by dividing it into the following three steps:

1. Setting up a full probability model—a joint probability distribution for all observable and unobservable quantities in a problem. The model should be consistent with knowledge about the underlying scientific problem and the data collection process.
2. Conditioning on the observed data: calculating and interpreting the appropriate posterior distribution—the conditional probability distribution of the unobserved quantities of ultimate interest, given the observed data.
3. Evaluating the fit of the model and the implications of the resulting posterior distribution: how well does the model fit the data, are the substantive conclusions reasonable, and how sensitive are the results to the modeling assumptions in step 1? In response, one can alter or expand the model and repeat the three steps.”

<sup>5</sup> Gelman et al [2013], page 3.

Jarrow and van Deventer go on to explain that the iterative process described above by Gelman et al is especially important in fitting Heath, Jarrow, and Morton parameters for the following reasons:

- a. Negative interest rates have been observed in Japan, Hong Kong, and many European countries, but many other countries, including the U.S. and Russia, have yet to report negative rates on central bank websites. In the U.S. case, the U.S. Department of the Treasury notes on its website that it overrides observed negative yields in the market with zero values.
- b. The “knowledge about the underlying scientific problem” from the historical data available is as follows: (1) negative rates are possible, (2) they are much less likely to occur than positive rates, (3) interest rate volatility that results when rates are negative is of high interest but the historical data is still limited or non-existent, depending on the country, and (4) an international data set would best shed light on this and other HJM issues.

There are other issues relevant to estimation:

- c. As noted by Heath, Jarrow, and Morton [1992], stochastic volatility driven by interest rate levels must be capped to avoid a positive probability of infinitely high interest rates.
- d. Subject to this cap, most market participants expect interest rate volatility to rise as rates rise and that the interest rate volatility that prevails when rates are negative represents the lowest level of volatility that would prevail. Historical experience with negative rates so far around the world makes it clear that interest rate volatility does not go to zero at any rate level.
- e. Most market participants believe that the empirical drift in forward rates that occurs (i.e., the change in observed empirical interest rates in the case where all interest rate shocks are zero) varies by the level of interest rates. The stochastic volatility model described in this paper assumes that empirical drift is a cubic function of annualized forward rates. The results show the expected mean reversion, which we discuss in Appendix B.

To summarize, a model validation effort in the Bayesian spirit would address at least these issues:

- Tests of smoothness of simulated curves
- Tests to confirm existence of negative rates in selected circumstances in the simulation
- Comparison of simulated risk neutral and empirical yields
- Time series distribution of simulated risk neutral and empirical yields
- Tests for the reasonableness of empirical drift in rates at each segment of the yield curve and stress tests of the rate factor volatility functions for reasonableness over the full range of rates that have been observed in the past. These tests are shown in Appendix B

We conduct an inspection of these issues with the aid of a forward-looking out-of-sample simulation of Government of Russia yields with the following specifications:

- Yield curve: Government of Russia yields
- Date of yields: Most recent date
- Number of scenarios: 500,000
- Simulation time horizon: Longest observable yield available (30 years in Russia)
- Simulation periodicity: 91 days (quarterly)

### **A. Smoothness Validation**

In the simulation, we select a random sample of 10 scenarios at each time step and visually examine forward rate graphics for smoothness. We can also use the discrete formula for smoothness given above to identify any outliers and examine the scenarios in question.

These graphs provide informal confirmation that nothing in the model estimation procedure has introduced artificial lumpiness in the simulated yield curves. A quantitative assessment of the smoothness of all 500,000 yield curves at each time step would provide the more formal confirmation that the yield curves simulated are realistically smooth.

## **B. Distribution of Simulated Risk Neutral and Empirical Rates**

We then examine the probability distributions of risk neutral and empirical simulated rates at various maturities over time. We seek to detect visually any points in time at which the simulated distribution of yields is strange or unrealistic.

We seek to determine that the simulation is reasonable from multiple dimensions. Rates can be negative but (for empirical yields) the probability of negative rates is low. On the long end of the maturity spectrum, we usually find that rates do rise to the 20% range but with a very low probability.

## **C. Time Series Distribution of Simulated Yields**

We also plot the time series graphs of the mean, median, high, low, and various percentiles for empirical rates. We again seek to determine that there are no unexpected variations in the distribution of empirical yields over time.

## **D. Simulation of the Term Premium**

The size of the “term premium” of actual zero-coupon yields over the expected level of the short rate is a topic of great interest to both academics and policy makers. In a stochastic volatility model, the term premium must be determined by simulation because in general there is no closed form solution for expected future rates. We again reach conclusions about whether the simulation produces results that are consistent with the “scientific knowledge” about the variation in interest rates around the world.

## **E. Probability of Negative Short-term Interest Rates**

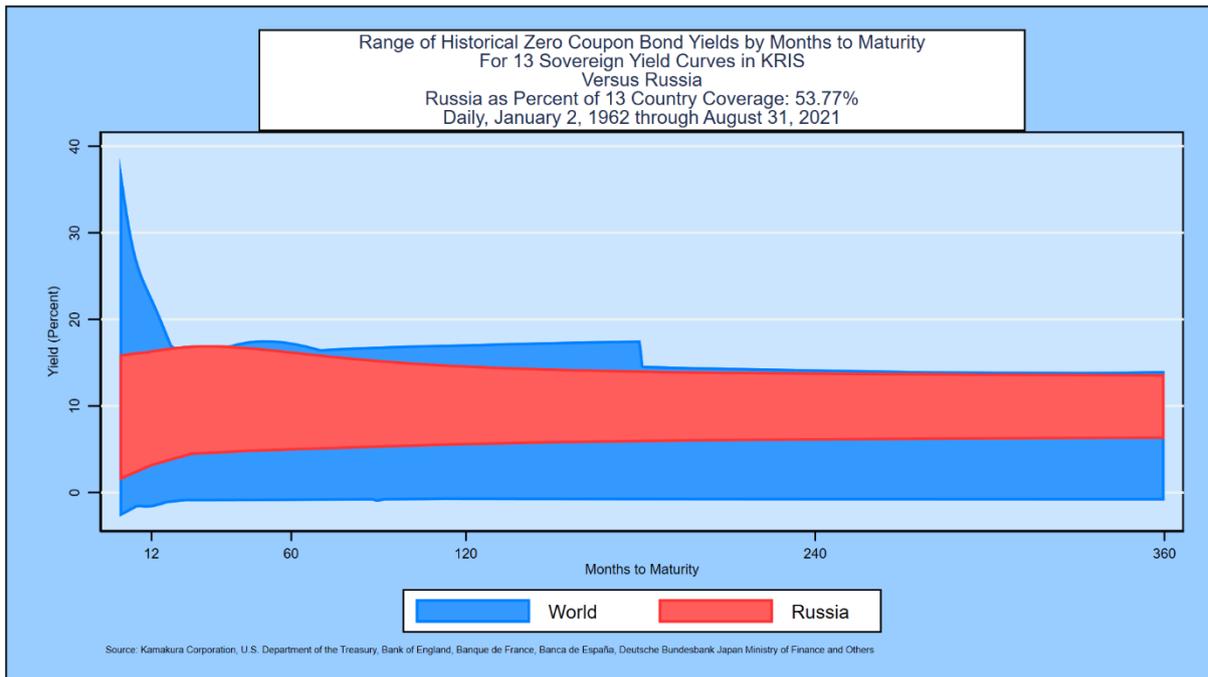
We also seek to determine whether or not there is a realistic possibility of negative rates in the simulation.

## **F. Comparing Russia Data with World Experience**

A common econometric challenge is fitting a model to the data at hand and measuring the risk that the model behaves badly out of sample. One measure of this risk for the Russia data set is to compare the range of interest rates experienced with the 13-country World data set over maturities out to 30 years. We use zero-coupon bond yields in a common format across countries and maturities for these calculations.

The following Exhibit XI shows that, measured by maturity range and interest rate range, Russia data spans 53.77% of World experience:

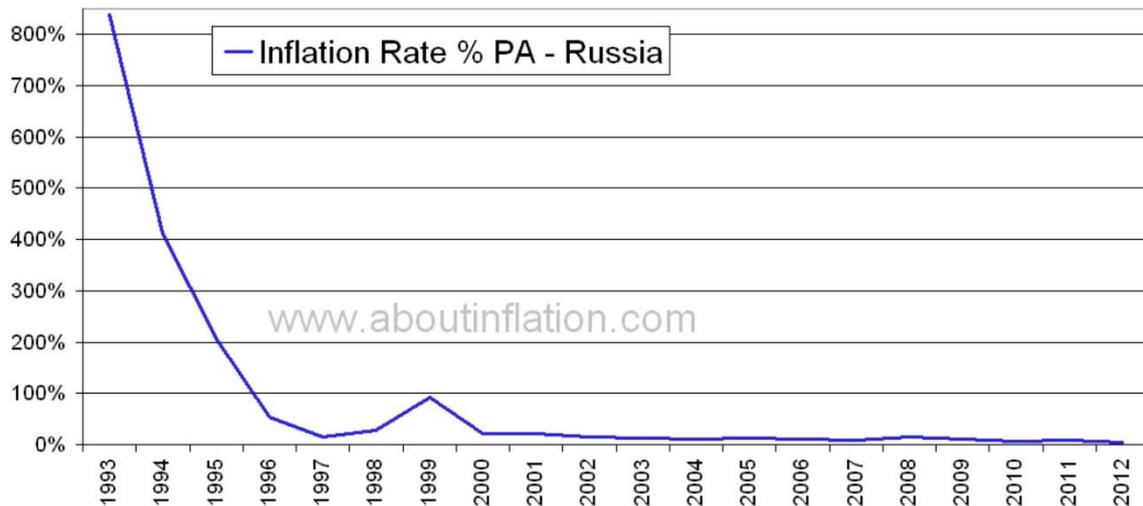
## Exhibit XI



Ironically, interest rate experience in the decade immediately prior to the Central Bank of Russia time series would have been immensely interesting, as this chart of the inflation rate shows:

Russia Inflation Rate Historical chart

Russia Inflation Rate Historical chart - 1993 to 2012



Source: [www.aboutinflation.com](http://www.aboutinflation.com)

Note the 1997 dip in Russia's inflation rate, just before the 1998 "black Thursday" spike. Banks and businesses became so used to the sky-inflation rates of the early 1990s that many of them found it hard to function at just a 20% inflation rate. In particular, the tactic of making a profit by delaying payments by a couple of months no longer worked; credit markets didn't work properly at that time. Russia's then-Prime Minister Viktor Chernomyrdin, infamous for his aphorisms and malapropisms, blurted

out this jab at Russia's monetarist central bankers: "Why do they want zero inflation? The only place with zero inflation is the cemetery!"

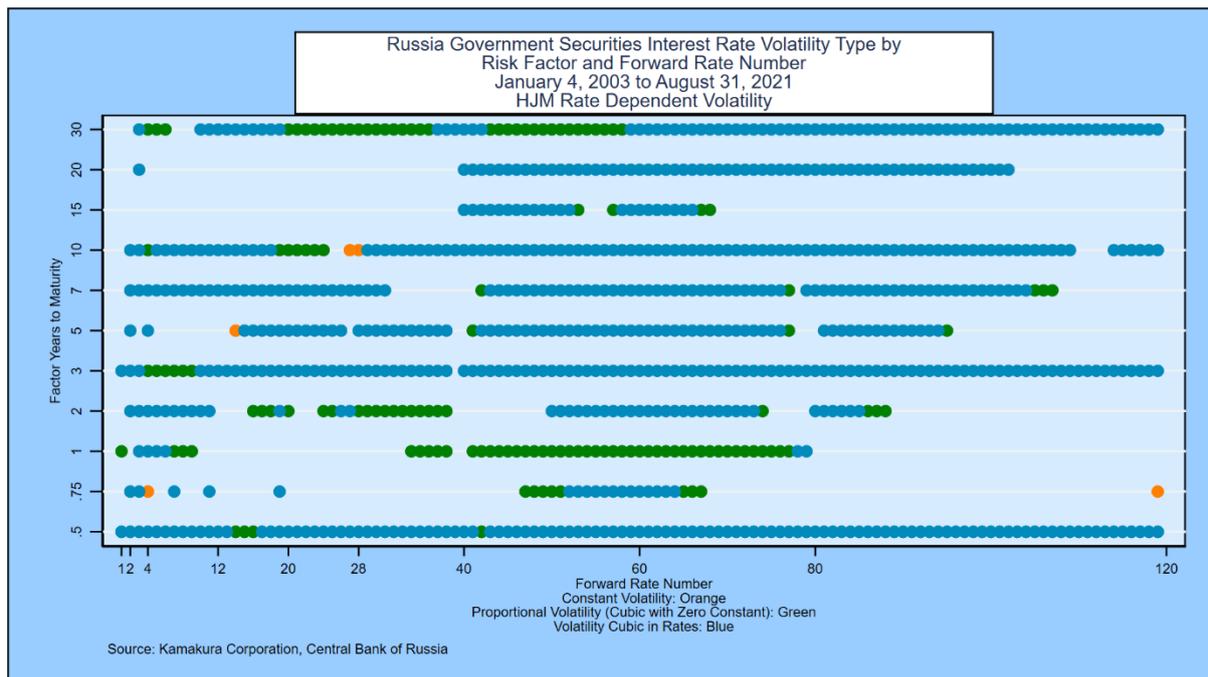
Additional stress tests for empirical drift of rates and for interest rate volatility reported in Appendix B confirm what this exhibit implies: for a realistic Government of Russia interest rate simulation that performs well "out of sample" (compared to the Russia yield history), the best term structure model to use would be the World model rooted in the experience of 13 countries, including Russia. This conclusion applies to all but one or two countries (the U.S. and the United Kingdom). Even in the U.S. case, the lack of experience with negative rates means the World model is likely to produce best results.

#### **IV. Conclusion**

The Government of Russia yield curve is driven by 11 factors, a number of factors that is near the average of what we have found for government yield curves in 12 other markets for which studies have been conducted. The January 4, 2003 to August 31, 2021 yield history for the Russia is fairly typical, but the lack of experience with very high and very low rates results in a sparser data environment than most analysts would want for a robust stand-alone single country term structure model.

The stochastic volatility assumption provides more accurate and more reasonable parameters than a constant volatility model, particularly in the context of Bayesian simulations as part of the model validation process. Exhibit XII summarizes the reasons for those conclusions:

## Exhibit XII: Statistical Significance Summary and Volatility Classification



The vertical axis lists the maturities used as risk factors by years to maturity of the underlying quarterly forward rate. The risk factors are the idiosyncratic movement of each of these forward rates. If the risk factor is statistically significant in explaining the movement of forward rates with the quarterly maturities listed on the horizontal axis, a dot is placed in the grid.

The nature of interest rate volatility for each combination of risk factor maturity and forward rate maturity is color coded. If the derived volatility is constant, the color code is orange. This is the affine specification. The graph shows immediately that only 5 of the risk factor maturity/forward rate maturity volatilities are consistent with the affine structure (the orange dots). The green and blue codes address the issue of whether interest rate volatility for that combination of risk factor maturity and forward rate maturity is zero or not when the forward rate is zero. If the measured volatility at a zero forward rate level is zero, the color code is green. Otherwise, the color code is blue. In both cases, the volatility is a stochastic function of the forward rates at the start of the simulation period.

The chart summarizes the fact that all 11 factors are statistically significant across the yield curve for Government of Russia bonds. The dominant derived interest rate volatility is the cubic stochastic volatility specification with a non-zero constant. An affine assumption for interest rate volatility is best fitting for a very small minority of the combinations of risk factor maturity and forward rate maturity.

## **Appendix A: Single-Factor versus Multiple-Factor Models**

Despite the overwhelming evidence across countries that government bond yields are driven by multiple factors, the use of single-factor term structure models in interest rate risk management systems remains common even in some of the world's largest financial institutions. This appendix asks and answers a number of important questions on the use of one-factor models that any sophisticated model audit would pose. Given the answers below, most analysts would conclude that one-factor term structure models are less accurate than a long list of multi-factor term structure models and that the one-factor models would therefore fail a model audit.

We address two classes of one-factor term structure models, all of which are special cases of the Heath, Jarrow, and Morton framework, in this appendix using data from the Government of Russia market and the World data set. Answers for other government bond markets cited in the references are nearly identical.

One-factor models with rate-dependent interest rate volatility:

Cox, Ingersoll, and Ross (1985)

Black, Derman and Toy (1990)

Black and Karasinski (1991)

One-factor models with constant interest rate volatility (affine models):

Vasicek (1977)

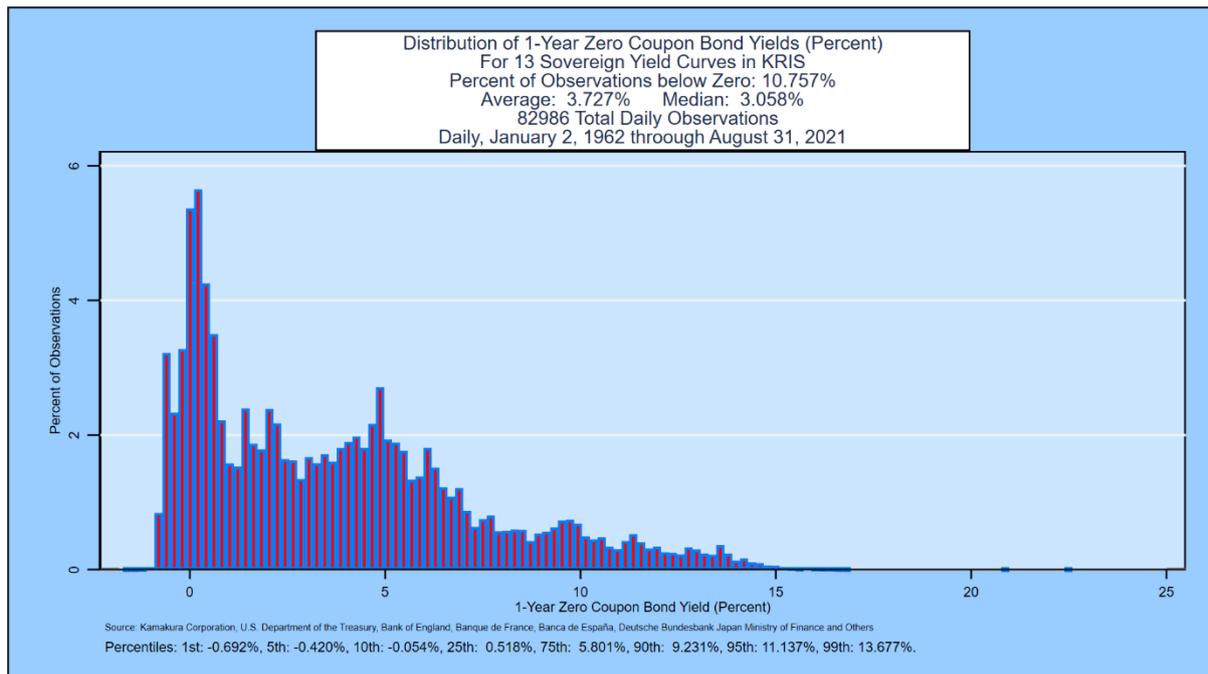
Ho and Lee (1986)

Extended Vasicek or Hull and White Model (1990, 1993)

### **Non-parametric test 1: Can interest rates be negative in the model?**

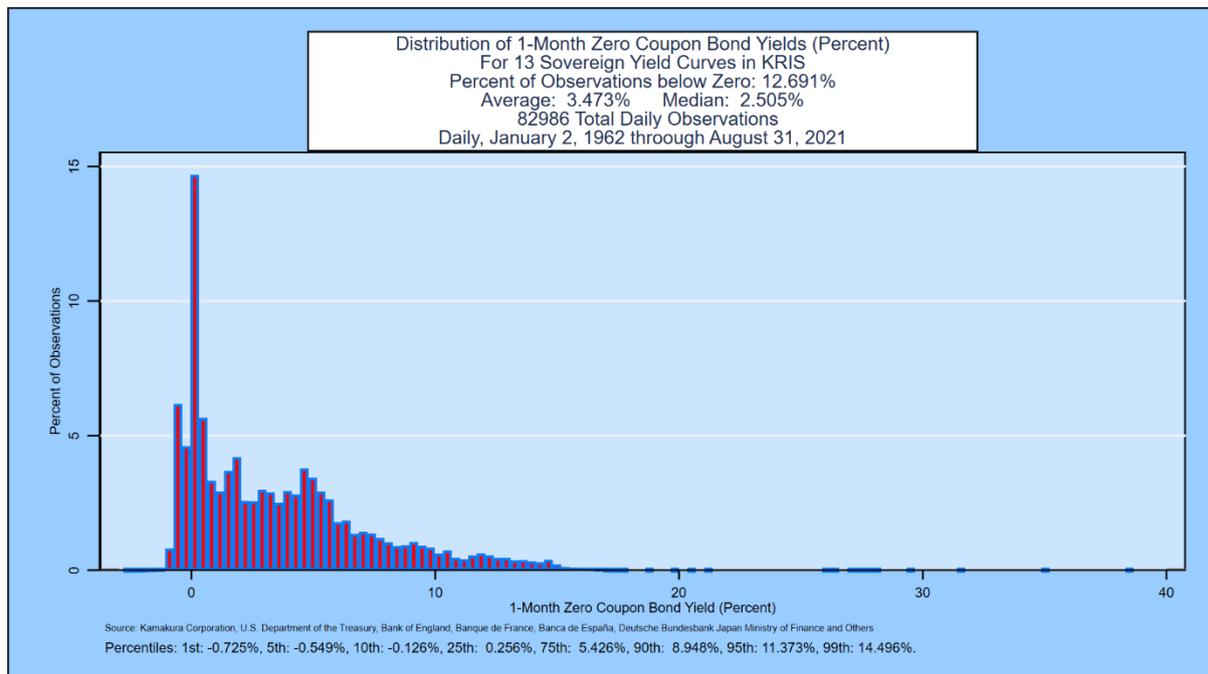
The one-factor models with rate-dependent interest rate volatility make it impossible for interest rates to be negative without ad hoc shifts in the simulated yield curves. Is this implication true or false? It is false, as this histogram of 1-year zero-coupon bond yields from government securities in 13 countries confirms. 10.757% of observations were negative:

## Exhibit A1



At a 1-month maturity, the percentage of negative rates is 12.691%:

## Exhibit A2



**Non-parametric test 2:** As commonly implemented, one-factor term structure models imply that all yields will either (a) rise, (b) fall, or (c) remain unchanged. This implication is false, as documented for the Government of Russia yield curve in Table

II. In fact, yield curves have twisted on 70.04% of the observation dates for the Government of Russia market.

**Non-parametric test 3:** The constant coefficient one-factor models imply that zero-coupon yields are normally distributed and so are the changes in zero-coupon yields. For all markets tested, this implication is rejected by three common statistical tests for all of the quarterly maturities for zero yields and for all of the quarterly changes. This histogram (shown above) of 1-month zero-coupon bond yields from government bond markets in 13 countries provides visual confirmation that normality is a poor approximation to the probability distribution of interest rates.

**Assertion A: There are no factors other than the short-term rate of interest that are statistically significant in explaining yield curve movements.** This assertion is false. Table VI shows, using principal components analysis, that 6-11 factors are needed to explain the movements of the Government of Russia yield curve. Exhibits IX and XII make the same point in more detail.

**Assertion B: There may be more than one factor, but the incremental explanatory power of the 2<sup>nd</sup> and other factors is so miniscule as to be useless.** This assertion is false, as the 2<sup>nd</sup> through 11<sup>th</sup> factors in the Government of Russia market explain 20.67% of forward rate movements, compared to 70.33% for the first factor alone. In most countries, the best “first factor” is not the short rate of interest used by many large banks; it is the parallel shift factor of the Ho and Lee model.

**Assertion C: A one-factor “regime shift” model is all that is necessary to match the explanatory power of the 2<sup>nd</sup> and other factors.** This assertion is also false. A recent study prepared for a major U.S. bank regulator also confirmed that a one-factor “regime shift” term structure model made essentially no incremental contribution toward resolving the persistent lack of accuracy in one-factor term structure models.

Finally, a Kamakura Corporation study comparing 1-factor and 10-factor Heath Jarrow and Morton simulation results shows that the simulated volatility of interest rates is underestimated by 61% to 83% in the 1-factor model. Moreover, the average level of simulated yields is biased lower. Finally, the probability of negative rates implied by the 1-factor model is also biased low as shown here using a U.S. Treasury simulation:

**Table A1****Kamakura Corporation****Comparison of Simulations for Different Term Structure Models****Base Model: U.S. Treasury Heath Jarrow and Morton 10-Factor Model, 1962-2020****Comparison Model: U.S. Treasury Heath Jarrow and Morton 1-Factor Model, 1962-2020****Starting Yield Curve: U.S. Treasury yields, March 5, 2021****Scenarios: 500,000****Time Horizon: 30 years****All Yields are Observable (Empirical) Yields in Percent****Copyright © Kamakura Corporation, 1990-2021, All Rights Reserved.**

<b>Zero Coupon Bond Yields</b>							
<b>Yield Maturity</b>	<b>Simulated Years Forward</b>	<b>Base Mean Yield</b>	<b>Comparison Mean Yield</b>	<b>Absolute Difference</b>	<b>Base Standard Deviation</b>	<b>Compare Standard Deviation</b>	<b>Percent Difference</b>
3 months	1	0.076	0.083	0.007	0.232	0.090	-61.2%
3 months	5	1.041	0.887	-0.154	1.003	0.247	-75.4%
3 months	10	1.588	1.114	-0.474	1.866	0.426	-77.2%
3 months	15	2.151	1.283	-0.868	2.580	0.580	-77.5%
5 years	1	1.244	1.238	-0.006	0.494	0.082	-83.4%
5 years	5	2.193	1.990	-0.203	1.190	0.254	-78.7%
5 years	10	2.815	2.299	-0.516	1.790	0.378	-78.9%
5 years	15	3.446	2.480	-0.966	2.349	0.452	-80.8%
10 years	1	1.922	1.913	-0.009	0.570	0.131	-77.0%
10 years	5	2.760	2.561	-0.199	1.160	0.271	-76.6%
10 years	10	3.468	2.897	-0.571	1.691	0.309	-81.7%
10 years	15	4.100	3.080	-1.020	2.136	0.343	-83.9%
<b>Par Coupon Bond Yields</b>							
3 months	1	0.076	0.083	0.007	0.232	0.090	-61.2%
3 months	5	1.044	0.888	-0.156	1.009	0.247	-75.5%
3 months	10	1.596	1.116	-0.480	1.887	0.427	-77.4%
3 months	15	2.165	1.286	-0.879	2.615	0.584	-77.7%
5 years	1	1.235	1.230	-0.005	0.488	0.082	-83.2%
5 years	5	2.189	1.984	-0.205	1.196	0.255	-78.7%
5 years	10	2.816	2.292	-0.524	1.821	0.382	-79.0%
5 years	15	3.459	2.475	-0.984	2.408	0.461	-80.9%
10 years	1	1.874	1.867	-0.007	0.551	0.124	-77.5%
10 years	5	2.711	2.513	-0.198	1.147	0.266	-76.8%
10 years	10	3.405	2.844	-0.561	1.695	0.316	-81.4%
10 years	15	4.401	3.024	-1.377	2.181	0.359	-83.5%

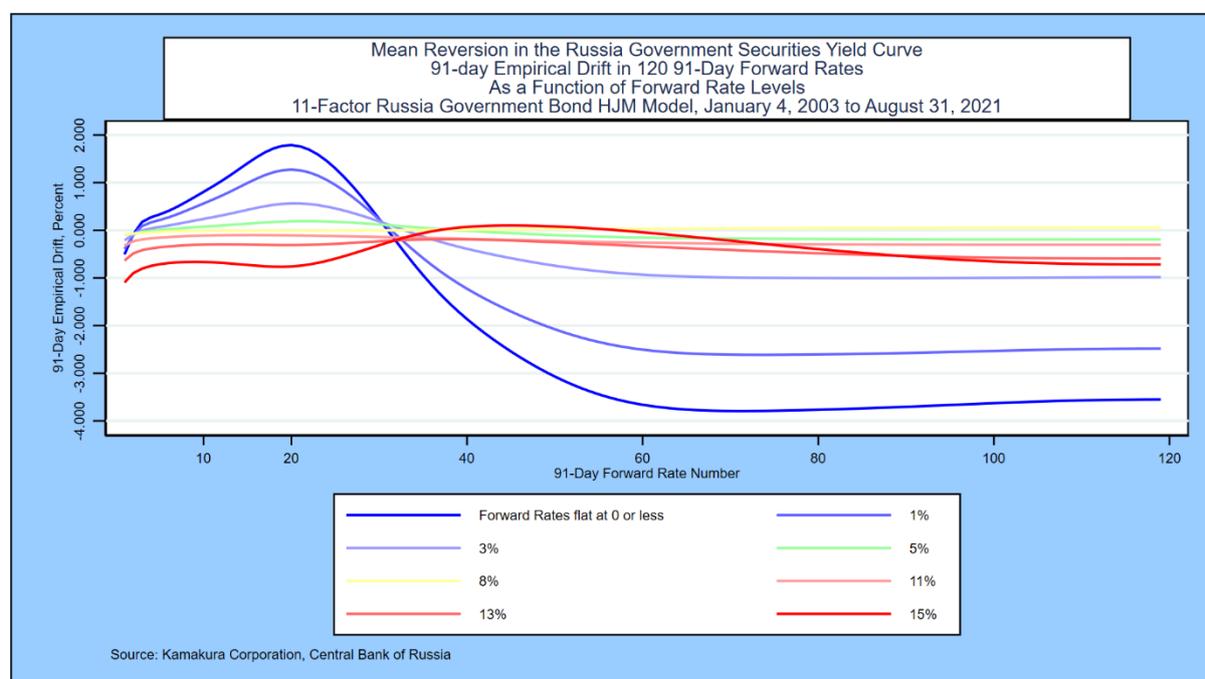
**Appendix B: Tests of Empirical Drift and Stress Tests of Rate Volatility**

Many of the Bayesian confirmations that out-of-sample simulations will be “consistent with our scientific knowledge of the problem” can in fact be gleaned without a full simulation. One practical modeling objective is to obtain out-of-sample rate scenarios that “do not blow up,” rising or falling to extreme values that are highly unlikely to occur in reality. Heath, Jarrow, and Morton make it clear how this objective is achieved in an HJM model:

1. First of all, rate shocks and rate volatilities are the same for the simulation of both risk-neutral and empirical rates.
2. This requires that the interest rate volatility functions be capped to prevent “explosions” in rate simulations, as we have done above. This is explained in the original Heath, Jarrow, and Morton [1992] paper.
3. HJM also make it clear that the risk-neutral drift in each segment of the yield curve is *completely determined* by the choice of volatility functions and the slope of the initial yield curve. This insight is not apparent to many practitioners because the early work of Vasicek [1977] and Hull and White 3 decades ago seems to embed mean reversion in risk-neutral interest rates, but in fact that is not the case.
4. The drift in empirical rates, where mean reversion is a testable hypothesis, is determined econometrically. In the model above, the drift function employed adds a cubic function of the corresponding forward rate to the base drift that applies when rates are zero or below.

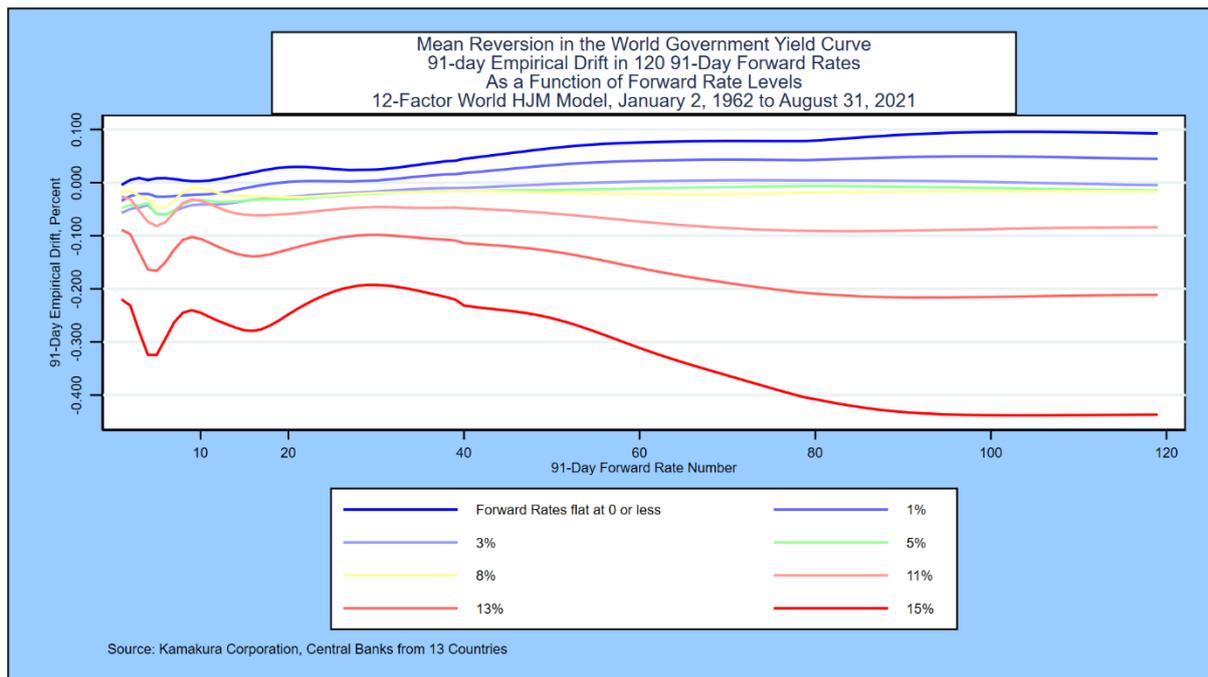
The Bayesian “scientific knowledge” of empirical rate movements leads us to expect some mean reversion in rates. When rates are high, we expect the empirical drift will be negative in each section of the empirical yield curve. When rates are low, we expect that the empirical drift to be positive. By applying hypothetical flat yield curves from zero (or less) to 15%, we can confirm that the Government of Russia HJM model is only modestly consistent with these expectations. The red lines represent the highest interest rate levels, and the empirical drift is negative across only the first part of the 30-year term structure for Russia Government Securities.

## Exhibit B1



The empirical drift for the World model through August 31, 2021 is shown in Exhibit B2.

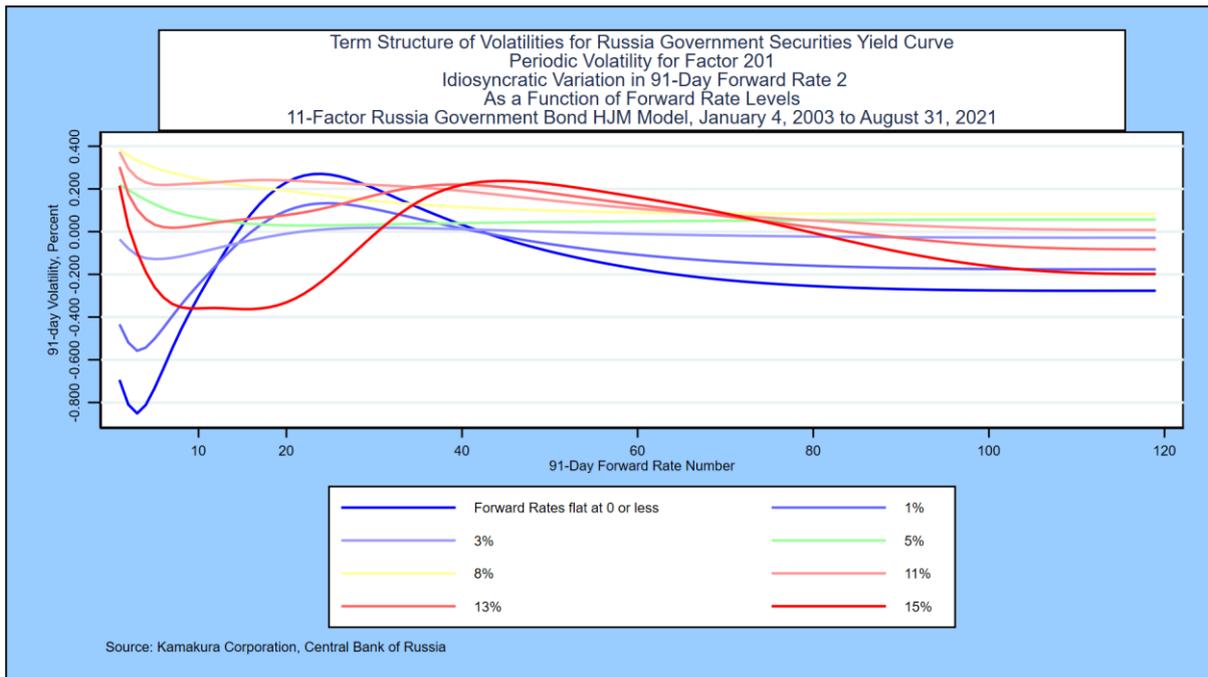
## Exhibit B2



Simply by visual inspection, it is obvious that the fitted drift in empirical rates for the World model shows a more robust version of mean reversion: rates drift down across the 30-year term structure when yields are high. Conversely, rates drift up when yields are low. From a Bayesian perspective, the World model's empirical drift is more consistent with our scientific knowledge of monetary economics and finance than the empirical drift based on a "Russia only" term structure model.

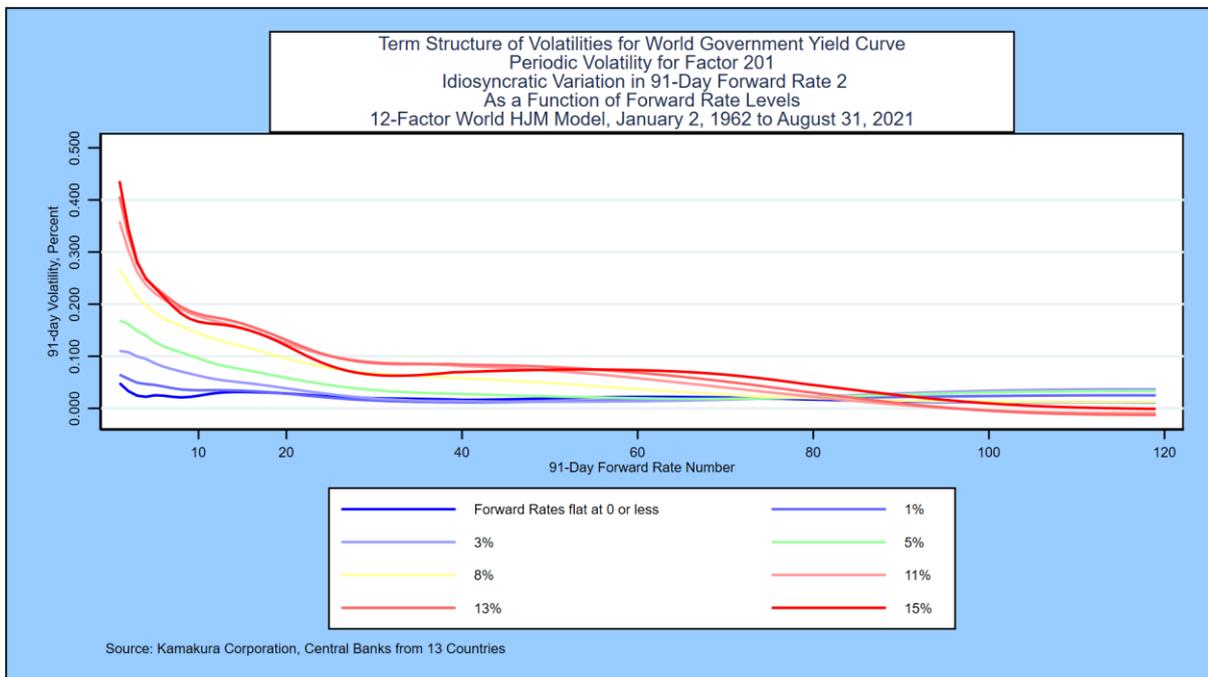
We can apply the same tests to the interest rate volatility functions to confirm their robustness and that they do not "blow up" at either very high or very low-rate levels. The first factor in the model above is the idiosyncratic movement in the first random forward rate segment. We expect this first factor to have volatility that declines with maturity and which is higher when rates are high. The graph below shows that those expectations are not that obvious in Government of Russia data:

### Exhibit B3



We again compare the Russia movements in interest rate volatility for the first factor with World experience, shown in the following graph:

### Exhibit B4

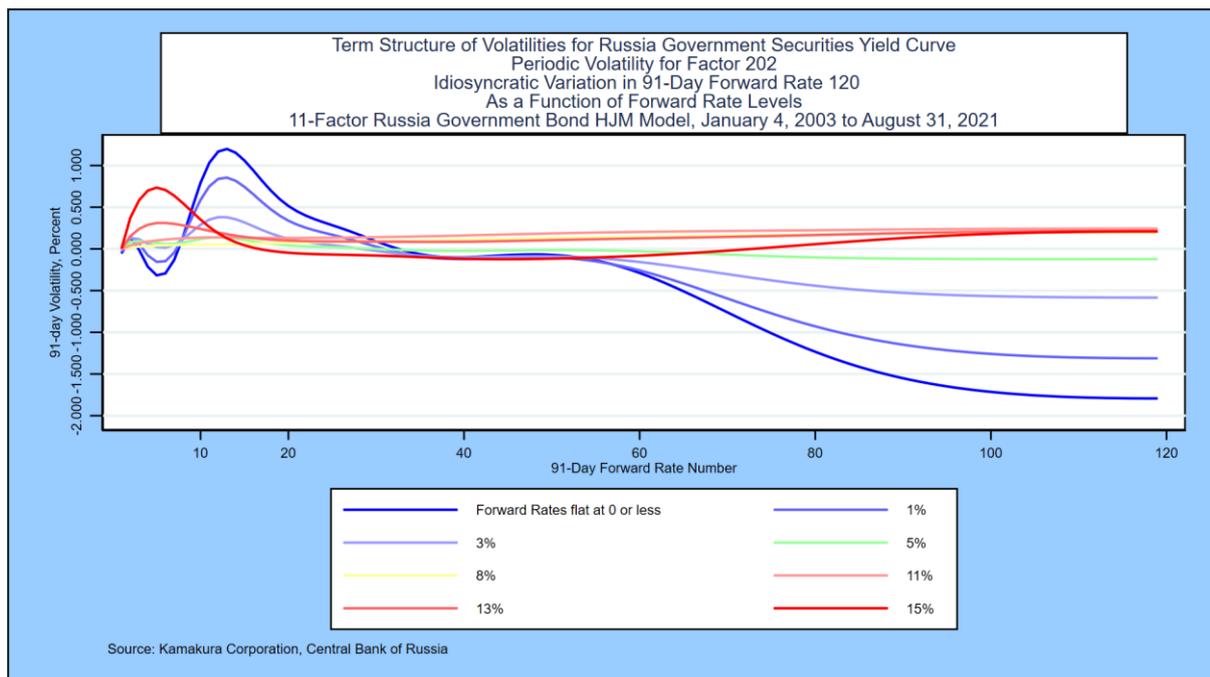


The World interest rate volatility shows a dramatic difference versus the Russia-only model. Since idiosyncratic movements in the short rate define the first factor, it makes sense that volatility of the short rate rises when rates are high and is closer to zero when rates are low. On the long end of the curve, volatility actually changes signs, pushing

the long end of the curve lower when the short rate spikes up. This is a volatility-based indicator of mean reversion. Again, we conclude that the World volatility function is more consistent with our knowledge of interest rate movements.

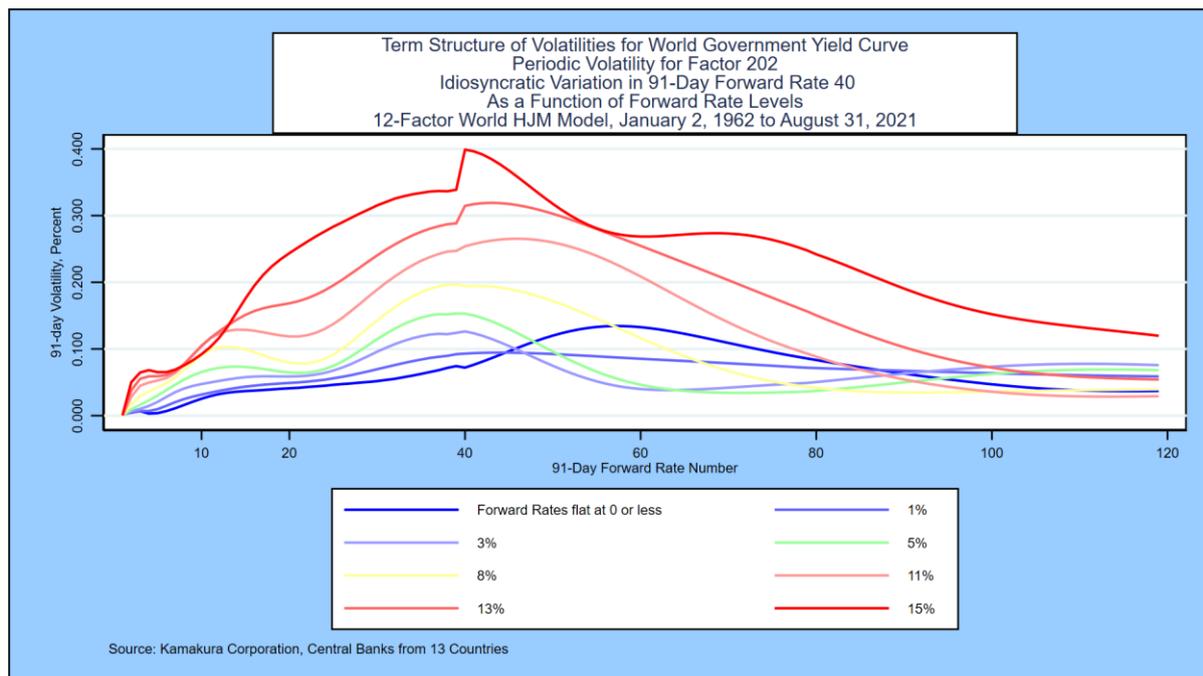
The second factor in the model is the idiosyncratic variation in the 91-day forward rate segment maturing in 30 years. Because this point is on the right-hand side of the yield curve in most countries, we expect volatility functions can be both positive and negative. Our main concern is that volatility does not blow up as we stress test the level of the forward rate curve. Again, our expectations are met, as shown below.

### Exhibit B5



How does this compare with volatility for the second factor (which is the idiosyncratic movements in the 91-day forward rate maturing in 10 years) for the World model? The answer is given in the graph below:

## Exhibit B6



Again, the results of the World model differ from the Russia-only model, but this is partly due to where the factor lies on the yield curve (at 30 years for Russia and 10 years for the World model). An upward shock in the 10-year forward rate drives up yields in the middle of the World yield curve, with smaller impacts on the short end and long end of the curve. Generally speaking, volatility increases as the level of the 10-year forward rate rises. Both models show realistic shifts in volatility but at different points on the yield curve.

A forward-looking simulation using both models from the same initial Government of Russia yield curve is another way to make these differences starkly apparent. Kamakura Corporation does such comparisons constantly and would be pleased to provide additional detail in this regard upon request to [info@kamakuraco.com](mailto:info@kamakuraco.com).

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